



The consequences of persistent inequality on social capital: A municipal-level analysis of blood donation data



Daniel Oto-Peralías^{a,*}, Diego Romero-Ávila^b

^a University of St Andrews, School of Management, The Gateway, North Haugh, St Andrews, Fife, KY16 9AJ, UK

^b Pablo de Olavide University, Department of Economics, Carretera de Utrera, Km. 1, Sevilla, Spain

ARTICLE INFO

Article history:

Received 20 October 2016

Accepted 25 November 2016

Available online 6 December 2016

JEL classification:

C21

O1

O18

R1

Keywords:

Persistent inequality

Land inequality

Culture

Social capital

Blood donation

ABSTRACT

This paper advances the hypothesis that persistent inequality affects cultural traits and undermines social capital. We use blood donation data at the local level in Southern Spain to document that, indeed, persistent inequality – as measured by land inequality – negatively affects blood donation, which indicates that it harms social capital. This evidence sheds new light into the debates on the consequences of inequality and the determinants of culture.

© 2016 Elsevier B.V. All rights reserved.

1. Introduction

This article addresses an unexplored consequence of inequality: its implications for culture and social capital formation. A society suffering severe inequality may develop social and political apathy. If the situation persists for a long time, this apathy becomes a cultural trait hindering the creation of social capital. By putting forward and testing this hypothesis we contribute to the debates on the consequences of long-term inequality (e.g., Engerman and Sokoloff, 2002; Easterly, 2007; Galor et al., 2009) and the causes of culture, and more particularly, of social capital (e.g., Nunn, 2012; Alesina and Giuliano, 2015; Guiso et al., 2016).

We focus on the effects of historical persistent inequality – in short *persistent inequality* –, which refers to a type of inequality that persists over a long historical period. This persistence makes inequality a feature of society that can create cultural traits. As part of a society's culture, social capital can be undermined by persistent inequality. Arguably, an unequal society that,

generations upon generations, excludes a large section of its population from having economic opportunities and acceptable living standards will hardly develop a strong sense of solidarity and commitment towards the common good. Once the cultural pattern has been created, it may persist even if society becomes more equal, thus constraining the potential for future economic growth.

We test this hypothesis using municipal-level data on blood donation in the Southern Spanish region of Andalusia (see Fig. 1). This is a case in point because Andalusia has suffered highly persistent inequality over its modern history, whose roots are largely exogenous (Oto-Peralías and Romero-Ávila, 2016, forthcoming). In addition, blood donation is collected in a centralized way by the Andalusian Health Service, which has provided us with a unique dataset on the number of donors per municipality. We find that land inequality – as a proxy for persistent inequality – has a non-negligible negative effect on blood donation, thus supporting the hypothesis that persistent inequality undermines social capital. Interestingly, it is historical inequality rather than current inequality which affects social capital, suggesting that the effect works through the creation of cultural patterns.

* Corresponding author.

E-mail addresses: dop2@st-andrews.ac.uk (D. Oto-Peralías), dromtor@upo.es (D. Romero-Ávila).



Fig. 1. Geographical context.

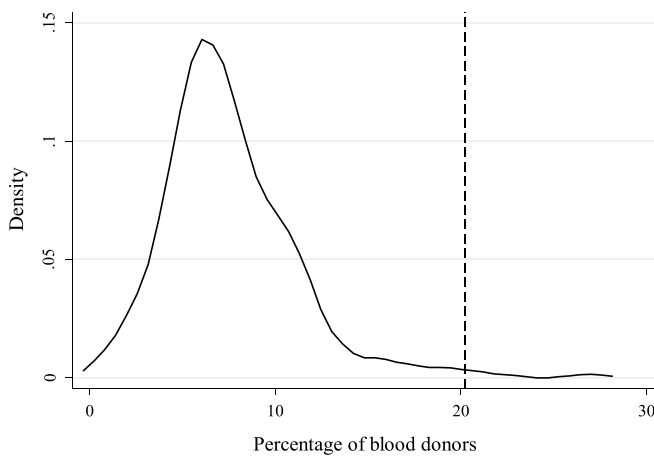


Fig. 2. Kernel density of the percentage of blood donors. Note: The vertical line reflects the value at which values are trimmed in the analysis (the 99th percentile).

2. Background and data

Following Guiso et al. (2011), social capital refers to “those persistent and shared beliefs and values that help a group overcome the free rider problem in the pursuit of socially valuable activities” (p. 419). We measure social capital through blood donation, which is widely considered to be a good proxy (e.g., Guiso et al., 2011; Nannicini et al., 2013). The Andalusian Health Service has provided us with a unique dataset containing the postal code of every person that has donated blood in any year of the period 2012–2014. Using this information and data on the population in age to donate blood (i.e., between 18 and 65), we create the variable “percentage of blood donors” at the municipal level. It is worth noting that blood donors’ data comes from an *ad hoc* extraction from administrative files conducted specifically for this study, and postal codes may contain errors. These errors are amplified in small municipalities, for which the denominator of the indicator is smaller. In addition, inhabitants of small municipalities have less access to blood collection units. Consequently, to reduce measurement errors we exclude municipalities with less than 1000 inhabitants in 2014, which only account for 1% of the Andalusian population. Fig. 2 depicts the Kernel density estimation of the distribution of the percentage of blood donors, which approximately follows a normal distribution slightly skewed to the right. To mitigate the influence of high values, we limit them to the 99th percentile (20.2%).

Our indicator of persistent inequality comes from the 1982 agricultural census and measures the percentage of utilized

agricultural area in holdings with 200 ha or more, computed considering only private agricultural holdings. This is a good proxy for historical persistent inequality because it reflects the incidence of *latifundia* in the municipality. Land concentration has been endemic in large parts of the Spanish and Andalusian geography, and can be traced back to the way land was colonized in the Middle Ages. Factors affecting land distribution in the remote past were contingent and exogenous to the development path of each territory (Oto-Peralías and Romero-Ávila, 2016, forthcoming). Fig. 3 provides some evidence on the persistence of land concentration throughout the 20th century in our sample of municipalities. Correlations are remarkably high, particularly bearing in mind that the indicators are not directly comparable. Tables A1 and A2 (Supplementary material, see Appendix A) provide the definitions and descriptive statistics of all variables used in the empirical analysis.

The persistent inequality suffered by municipalities characterized by high land concentration may have shaped the local culture and undermined social capital. The miserable living conditions of landless workers in Andalusia were pretty much those of a marginalized social group. Olavide (1768–1996), Intendant of Seville in the second half of the 18th century, described them as “the unhappiest men that I know in Europe [...] half-a-year laborers, and the other half beggars”. This situation of inequality and dependence on the landowner persisted well into the 20th century (Oto-Peralías and Romero-Ávila, 2016). Arguably, a society that over a long period of time has suffered such high levels of inequality will not develop a sense of commitment towards the public good and solidarity among its members. Social groups such as landless workers that have been oppressed, marginalized, or simply placed in a systematic inferior position, are very unlikely to view the state (or the society as a whole) as representing or sharing their own interests. This feeling of apathy, generations upon generations, leads to a culture of lack of trust, cooperation, or more broadly, social capital.

Finally, an interesting aspect of land inequality is that it was a very important determinant of overall economic inequality in the past, while its current relevance is much lower since agriculture only employs today a small fraction of the population (8.4% in Andalusia in 2014 according to the Labor Force Survey). Thus, in historical and cultural terms the effect of land inequality can be interpreted as a cumulative effect over time, which creates cultural patterns.

3. Empirical results

3.1. Baseline results

Column 1 of Table 1 reports the bivariate regression of the percentage of blood donors in 2014 on land concentration in 1982. The coefficient is negative and highly statistically significant, which provides initial support for our hypothesis. Column 2 adds two demographic factors that are relevant to explain blood donation, namely, total population and population’s average age (in linear and quadratic terms), while column 3 includes geographic indicators such as a coast dummy, distance to the capital city (linear and squared), and a capital city dummy. In both cases the coefficient on land concentration remains negative and statistically significant. Column 4 adds additional geographic controls that may influence both blood donation and land concentration (altitude, ruggedness, soil quality, rainfall, and average temperature). This is our baseline specification. The coefficient on land inequality is large and precisely estimated. Going from a municipality without large estates (0% in land concentration) to a municipality with a high concentration of land

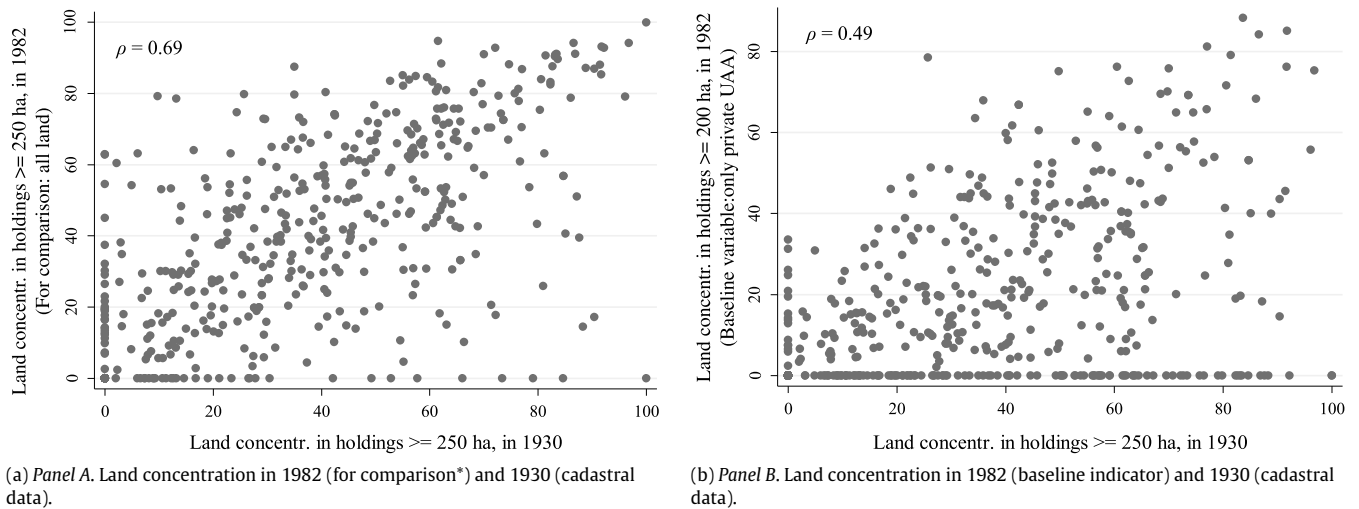


Fig. 3. Persistence in land concentration. Note: *Land concentration in 1982 in Panel A has been created for comparison purposes with the measure of 1930, which considers all types of land (see more details in Table A1).

Table 1
The effect of land concentration on blood donation: Baseline results.

Dependent variable is the percentage of blood donors						
	(1)	(2)	(3)	(4)	(5)	(6)
Land concentration in 1982	-0.017*** (0.006)	-0.014** (0.006)	-0.022*** (0.006)	-0.022*** (0.007)		
Land concentration in 1962					-0.015*** (0.005)	
Land concentration in 1930						-0.02*** (0.005)
Population		-0.04*** (0.006)	-0.027*** (0.006)	-0.025*** (0.006)	-0.026*** (0.006)	-0.032*** (0.007)
Population squared		56.945*** (10.756)	33.427*** (8.885)	31.325*** (9.022)	32.554*** (8.998)	57.143*** (10.144)
Population's average age		3.256*** (0.622)	2.544*** (0.615)	2.112*** (0.613)	2.005*** (0.628)	1.601*** (0.707)
Population's average age squared		-0.036*** (0.008)	-0.029*** (0.007)	-0.024*** (0.007)	-0.023*** (0.008)	-0.018*** (0.009)
Coast dummy			-2.708*** (0.278)	-2.27*** (0.366)	-2.24*** (0.381)	-2.417*** (0.414)
Distance to capital city			4.516** (1.792)	5.487*** (1.988)	4.965** (2.002)	5.793** (2.28)
Distance to capital city squared			-2.464 (1.583)	-3.554* (1.773)	-3.22* (1.775)	-4.123** (1.969)
Capital city dummy			2.834* (1.356)	3.132* (1.413)	3.005** (1.463)	2.197* (1.272)
Altitude				0.002* (0.001)	0.002* (0.001)	0.002 (0.001)
Ruggedness				-0.003* (0.002)	-0.002 (0.002)	-0.001 (0.002)
Soil quality				0.028 (0.25)	-0.091 (0.266)	0.028 (0.283)
Rainfall				0.149* (0.085)	0.13 (0.083)	0.061 (0.115)
Average temperature				0.09 (0.187)	0.077 (0.188)	0.14 (0.211)
R-squared	0.01	0.17	0.23	0.24	0.23	0.24
Obs.	570	570	570	570	569	509

Notes: Variables descriptions are provided in Table A1 (Supplementary material, see Appendix A). All regressions are estimated by OLS. The estimations include a constant term, which is omitted for space considerations. Robust standard errors are in parentheses.

* Denotes statistical significance at the 10% level.
 ** Denotes statistical significance at the 5% level.
 *** Denotes statistical significance at the 1% level.

(90%) decreases blood donors by almost 2 percentage points (i.e., $(90-0) * (-0.022) = -1.98$).

The next two columns show that the result does not depend on a specific variable of land concentration. Column 5 employs the percentage of land in holdings equal to or greater than 200 ha from the 1962 agricultural census, while column 6 uses the percentage of land in holdings greater than 250 ha from cadastral data in 1930. The results obtained with these alternative measures are very similar to the baseline findings. This is as expected since land distribution has remained largely unchanged over time.

3.2. A two-stage least squared (2SLS) framework

Following our previous work (Oto-Peralías and Romero-Ávila, 2016, forthcoming), we argue that land inequality can be largely considered as an exogenous factor due to the way the territory was colonized in the Middle Ages after the Christian conquest. This fact makes us confident that the uncovered negative effect of persistent inequality on social capital is not driven by reverse causality or omitted variables. To further pursue this issue, this section explicitly exploits the fact that Andalusia was divided during the Middle Ages by a frontier that ceased to exist in 1492 when the Catholic Monarchs conquered the Nasrid Kingdom of Granada (see Fig. 1). The Castilian part of Andalusia was colonized under the conditions of an insecure frontier region, which led to the concentration of land in the hands of the military elite, i.e., the nobility. In contrast, once conquered, the former Nasrid Kingdom of Granada was colonized under very different premises, that is, as a secure region, with land distribution evolving in a relatively more equal way (see details in Oto-Peralías and Romero-Ávila, forthcoming). Since the territory close to the frontier is geographically very similar (as shown in Table 2 in Oto-Peralías and Romero-Ávila, forthcoming), we can use this historical accident as a source of exogenous variation in persistent inequality within a 2SLS framework.

Columns 1 and 2 of Table 2 employ the whole sample, while columns 3 and 4 focus on municipalities within 25 km of the frontier. The instrument is a dummy variable capturing whether the municipality belonged to the Castilian part of Andalusia in the Middle Ages. This variable exerts a strong positive effect on land concentration, thereby indicating that it is a relevant instrument. The second stage, which uses only that part of land concentration that is due to the frontier dummy, reports a large, negative, and highly significant coefficient. The increase in the size of the effect suggests that the OLS estimates are downward biased. Moreover, it could suggest that the exclusion restriction does not hold. In this regard, one possible interpretation is that, as documented in Oto-Peralías and Romero-Ávila (forthcoming), the frontier of Granada also affected the concentration of political power by the nobility. Thus, if we consider the second-stage coefficient on land concentration to be capturing inequality in a broad sense (both in economic and political terms), this would also justify its larger effect. Notwithstanding, due to lack of data on other potential channels, it is impossible to test whether the frontier has only affected social capital through persistent inequality. Therefore, we cannot assure that persistent inequality is the only channel through which the presence of the frontier of Granada affected social capital, but it is likely to be the dominant one. All in all, this 2SLS exercise provides evidence consistent with the hypothesis that persistent inequality undermines social capital.

3.3. Additional robustness checks

Our reading of the results is that persistent inequality – proxied by land concentration – contributes to create a culture of low cooperation and low commitment towards the common

good, which means low social capital. Therefore, this implies that persistent inequality has direct cultural implications. An alternative interpretation would be that persistent inequality harms economic development in the long-run, with poorer communities ending up with less social capital. This would imply an indirect effect working through economic development. We try to address this issue by adding proxies for income to our baseline model. Columns 1–4 of Table 3 control for the percentage of population with secondary and higher education, the average number of vehicles per household, average socio-economic condition, and average gross income. All these variables enter with a positive coefficient, which is statistically significant only in column 3. Reassuringly, the coefficient on land concentration remains unchanged.

One may also wonder whether it is current inequality, rather than historical inequality, what matters. Current inequality can be measured through income inequality in 2007. Data on this variable at the local level is seldom available. For the Spanish case, there are data available for municipalities larger than 5000 inhabitants (Hortas-Rico and Onrubia, 2014). Column 5 estimates our baseline model with this smaller sample of municipalities. The coefficient on land concentration is larger, which is likely due to the fact that measurement errors in the dependent variable diminish as municipality size increases. Column 6 includes an income Gini index, which enters with a negative but statistically insignificant coefficient. Column 7 includes both inequality variables. Remarkably, the coefficient on persistent inequality (i.e., land concentration) remains negative and significant, while that on current inequality is again insignificant.

Finally, our findings are also robust to: (i) correcting standard errors for spatial dependence; (ii) trimming values of the percentage of blood donors at the 95th percentile (rather than at the 99th); (iii) applying different cutoffs of municipalities' population; and (iv) using electoral turnout as an alternative indicator of social capital (see Supplementary material for details).

4. Conclusions

This paper advances and tests the hypothesis that persistent inequality undermines social capital. Using data on blood donors as a proxy for social capital, and land concentration as a measure of persistent inequality, our municipal-level analysis finds support for this hypothesis. Land inequality exerts a non-negligible negative effect on the percentage of blood donors, which is robust to the inclusion of a wide array of demographic and geographic controls. The effect is also robust to controlling for several proxies for income, which suggests that our findings are not driven by the fact that inequality negatively affects economic growth. In addition, we find that what matters is historical inequality rather than current (income) inequality. This is consistent with the fact that once the cultural trait is created, it may persist even if society becomes more equal. In all, the evidence provided sheds new light into the debates on the consequences of inequality and the determinants of culture.

Acknowledgments

We are indebted to the Andalusian Health Service (*Centro Regional de Transfusión Sanguínea de Sevilla* and *Dirección General de Asistencia Sanitaria y Resultados en Salud*) for providing us with data on blood donors. Ivan Dondukov provided excellent research assistance. Financial support by the Carnegie Trust for the Universities of Scotland through grant 70433 and the Spanish Ministry of Economics and Competitiveness through grant ECO2012-35430 is gratefully acknowledged.

Table 2

The effect of land concentration on blood donation: A 2SLS exercise.

	Dependent variable in the second stage is the percentage of blood donors			
	Whole sample		Within 25 km of the Frontier	
	First stage	Second stage	First stage	Second stage
	(1)	(2)	(3)	(4)
Land concentration in 1982		−0.109 ^{***} (0.038)		−0.164 ^{**} (0.081)
Castilian part of Andalusia	11.744 ^{***} (1.951)		8.802 ^{***} (2.839)	
Demographic and geographic controls	Yes	Yes	Yes	Yes
F-stat of the instrument	36.220		9.613	
Partial R-squared	0.050		0.060	
R-squared	0.219		0.292	
Obs.	570	570	167	167

Notes: Variables descriptions are provided in Table A1 (Supplementary material, see Appendix A). All regressions are estimated by 2SLS. The estimations include a constant term, which is omitted for space considerations. Robust standard errors are in parentheses.

* Denotes statistical significance at the 10% level.

** Denotes statistical significance at the 5% level.

*** Denotes statistical significance at the 1% level.

Table 3

Robustness to alternative interpretations.

	Dependent variable is the percentage of blood donors						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Land concentration in 1982	−0.022 ^{***} (0.007)	−0.022 ^{***} (0.007)	−0.020 ^{***} (0.007)	−0.022 ^{***} (0.007)	−0.028 ^{***} (0.007)		−0.029 ^{***} (0.007)
Education level (2001)	0.009 (0.016)						
Average number of vehicles per household (2001)		0.15 (1.155)					
Average socio-economic condition (2001)			3.028 [*] (1.751)				
Average gross income (2013)				0.79 (0.703)			
Income Gini index (2007)						−1.367 (2.771)	−3.034 (2.578)
Demographic and geographic controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes
R-squared	0.240	0.24	0.240	0.240	0.400	0.370	0.41
Obs.	569	569	569	570	234	236	234

Notes: Variables descriptions are provided in Table A1 (Supplementary material, see Appendix A). All regressions are estimated by OLS. The estimations include a constant term, which is omitted for space considerations. Robust standard errors are in parentheses.

* Denotes statistical significance at the 10% level. ** Denotes statistical significance at the 5% level.

*** Denotes statistical significance at the 1% level.

Appendix A. Supplementary material

Supplementary material related to this article can be found online at <http://dx.doi.org/10.1016/j.econlet.2016.11.037>.

References

- Alesina, A., Giuliano, P., 2015. Culture and institutions. *J. Econom. Lit.* 53, 898–944.
- Easterly, W., 2007. Inequality Does Cause Underdevelopment: Insights from a new instrument. *J. Dev. Econ.* 84, 755–776.
- Engerman, S.L., Sokoloff, K.L., 2002. Factor endowments, inequality, and paths of development among new world economies. *Economia* 3, 41–109.
- Galor, O., Moav, O., Vollrath, D., 2009. Inequality in land ownership, the emergence of human capital promoting institutions and the great divergence. *Rev. Econom. Stud.* 76, 143–179.
- Guiso, L., Sapienza, P., Zingales, L., 2011. Civic capital as the missing link. In: Benhabib, J., Bisin, A., Jackson, M. (Eds.), *Social Economics Handbook*. North-Holland.
- Guiso, L., Sapienza, P., Zingales, L., 2016. Long-term persistence. *JEEA* <http://dx.doi.org/10.1111/jeea.12177>.
- Hortas-Rico, M., Onrubia, J., 2014. Renta personal de los municipios españoles y su distribución: metodología de estimación a partir de microdatos tributarios, Estudios sobre la Economía Española 2014–12, FEDEA.
- Nannicini, T., Stella, A., Tabellini, G., Troiano, U., 2013. Social capital and political accountability. *Amer. Econ. J. Econ. Policy* 5 (2), 222–250.
- Nunn, N., 2012. Culture and the historical process. *Econ. Hist. Dev. Reg.* 27 (S1), 108–126.
- Olavide, Pablo de, 1768–1996. Informe sobre la Ley Agraria. In: Merchán, A. (Ed.), *La Reforma Agraria en Andalucía: el Primer Proyecto Legislativo*. Universidad de Sevilla, Sevilla.
- Oto-Peralías, D., Romero-Ávila, D., 2016. The economic consequences of the Spanish Reconquest: The long-term effects of medieval conquest and colonization. *J. Econ. Growth* 21, 409–464. <http://dx.doi.org/10.1007/s10887-016-9132-9>.
- Oto-Peralías, D., Romero-Ávila, D., 2016. Historical frontiers and the rise of inequality. The case of the frontier of Granada. *JEEA* <http://dx.doi.org/10.1093/jeea/jvw004> (forthcoming).